#### HERTENTAMEN

# Kansrekening en Statistiek -4I

#### Wednesday 31 August 2005

#### Instructions

- Write your name and studentnumber on every page
- Motivate all answers clearly
- Use of a calculator will not be necessary and is not allowed
- Use of Rice's book is allowed
- If you finish the exam early, hand in your work and leave quietly
- Do not attempt to cheat or confer with your fellow students

```
Points per question (Grade = total/8 + 1)
```

```
      1(a):
      5
      2(a):
      5
      3(a):
      5

      1(b):
      5
      2(b):
      5
      3(b):
      5

      1(c):
      5
      2(c):
      6
      3(c):
      * 5

      1(d):
      6
      2(d):
      5
      3(d):
      5

      3(e):
      5

      3(f):
      5
```

## Good luck!

[DO NOT TURN THIS PAGE UNTIL SO INSTRUCTED]

#### PROBLEM 1

Let X in  $\mathbb{R}$  be continuously distributed with probability density function  $f_X : \mathbb{R} \to \mathbb{R}$  of the form:

$$f_X(x) = \begin{cases} Cx^2(2-x)^2, & \text{if } 0 \le x \le 2, \\ 0, & \text{if } x < 0 \text{ or } x > 2. \end{cases}$$

- a. (5 points)

  Calculate the normalization constant C.
- b. (5 points)Show that E(X) = 1.
- c. (5 points) Show that the second moment  $E(X^2)$  satisfies

$$E(X^2) = \frac{128}{105}C,$$

and, using your answer under part a., show that Var(X) = 8/7.

d. (6 points)

Give the distribution function  $F_X : \mathbb{R} \to [0,1]$  of X.

Based on your answer, demonstrate that the median of X equals 1.

#### PROBLEM 2

Let X, Y be two continuous random variables, jointly distributed according to a joint probability density function  $f_{XY}: \mathbb{R}^2 \to \mathbb{R}$ ,

$$f_{XY}(x,y) = \begin{cases} Ce^{-\lambda(x+y)}, & \text{if } x \ge 0 \text{ and } y \ge 0, \\ Ce^{-\lambda(x-y)}, & \text{if } x \ge 0 \text{ and } y < 0, \\ Ce^{-\lambda(-x+y)}, & \text{if } x < 0 \text{ and } y \ge 0, \\ Ce^{\lambda(x+y)}, & \text{if } x < 0 \text{ and } y < 0. \end{cases}$$

depending on a parameter  $\lambda > 0$ .

a. (5 points)

Calculate the normalization constant C

(NB: your answer will be an expression involving  $\lambda$ ).

b. (5 points)

Show that the marginal probability density function  $f_X : \mathbb{R} \to \mathbb{R}$  for X has the following form:

$$f_X(x) = egin{cases} rac{\lambda}{2} \, e^{-\lambda \, x}, & ext{if } x \geq 0, \ rac{\lambda}{2} \, e^{\lambda \, x}, & ext{if } x < 0. \end{cases}$$

c. (6 points)

Show that the random variables X and Y are independent (motivating your answer by an equality of probability density functions).

What relation does the independence of X and Y provide between Var(X+Y), Var(X) and Var(Y)?

d. (5 points)

Show that:

$$Var(X+Y) = \frac{4}{\lambda^2}.$$

(NB: You may use that  $\int_0^\infty u^n e^{-u} du = n!$ .)

## PROBLEM 3

Consider an experiment in which we measure X, a random variable with a distribution depending on an unknown parameter  $\kappa > 0$ . The probability density function  $f_{\kappa} : \mathbb{R} \to \mathbb{R}$  for X is given by:

$$f_{\kappa}(x) = \begin{cases} 0, & \text{if } x < 0, \\ \frac{1}{\kappa\sqrt{\pi}} x^{-1/2} e^{-\frac{x}{\kappa^2}}, & \text{if } x \ge 0. \end{cases}$$

Assume that we have an *i.i.d.* sample  $X_1, X_2, \ldots, X_n$  of repeated measurements of X. Parts a and b of this problem concern the moment estimator for  $\kappa$ .

a. (5 points)

By partial integration, show that:

$$\int_0^\infty x^{1/2} e^{-x/\kappa^2} dx = \frac{\kappa^2}{2} \int_0^\infty x^{-1/2} e^{-x/\kappa^2} dx.$$

Use this equality to show that  $E(X) = \kappa^2/2$ .

b. (5 points)

Give the first moment equation and derive the form of the moment estimator  $\tilde{\kappa}$  for the parameter  $\kappa$ .

Indicate clearly in which step of your derivation the restriction  $\kappa > 0$  plays a role.

Next, we turn to the maximum-likelihood estimator of  $\kappa$ . As it turns out, it is conventient to use, instead of  $\kappa$  itself, its square  $\theta = \kappa^2$  as a parameter. This does not change the probability density functions in the model, only the parametrization. To obtain the maximum-likelihood estimator for  $\kappa$ , we shall solve the maximum-likelihood problem for  $\theta$  first and then derive the maximum-likelihood estimator for  $\kappa$ .

### c. (5 points)

What is the domain of the new parameter  $\theta$ ?

Give the form for the probability density function  $f_{\theta} : \mathbb{R} \to \mathbb{R}$  parametrized by the parameter  $\theta$ .

Also give the log-likelihood  $\ell(\theta)$  based on the *i.i.d.* sample  $X_1, \ldots, X_n$ .

## d. (5 points)

Write down the first derivative  $\dot{\ell}(\theta)$  of  $\ell(\theta)$  with respect to  $\theta$ .

Solve the equation  $\ell(\theta) = 0$  to show that the maximum-likelihood estimator  $\hat{\theta}$  for  $\theta$  equals:

$$\hat{\theta}(X_1,\ldots,X_n)=2\bar{X}.$$

## e. (5 points)

Based on part d., give the maximum-likelihood estimator  $\hat{\kappa}$  for  $\kappa$ .

Explain your answer clearly.

(Hint: this does not require any calculation: extrema of a function are parametrization-invariant.)

#### f. (5 points)

Is the maximum-likelihood estimator  $\hat{\theta}$  for  $\theta$  biased or unbiased? Can we conclude from this that the maximum-likelihood estimator  $\hat{\kappa}$  for  $\kappa$  is baised/unbiased?