VU University	Statistical Data Analysis
Faculty of Sciences	28 May 2014

## **SOLUTION**

# PART 1

## Question 1 [6 points]

- a. [2 points] Location: minimum  $\approx 0$ , maximum  $\approx 8.8$ , median  $\approx 1.4$ , IQR  $\approx 2.2$ , data are skewed to the right, and the sample contains two high extremes (around 7.0 and 8.8).
- b. [2 points] The QQ-plot against the exponential distribution shows the most straight line, and therefore that location-scale family is most appropriate.
- c. [2 point] Fitting a line by hand in the QQ-plot yields approximately y = 0 + 2x, so location a = 0 and scale b = 2. Since we chose the exponential distribution this corresponds to the Exp(1/2) distribution.

## Question 2 [5 points]

- a. [1 point] The given expression is correct.
- b. [2 points] Under the null hypothesis the test statistic has approximately a  $\chi^2$ -distribution with k-1 degrees of freedom. This approximation is reliable if  $np_i \geq 5$  for  $i=1,\ldots,k$ . In other words: the expected number of observations under the null hypothesis has to be at least 5 in all intervals.
- c. [2 point] One could you the one sample Kolmogorov–Smirnov test, which is also a goodness-of-fit test for a simple null hypothesis.

## Question 3 [7 points]

- a. [3 points] Given a sample  $Z_1, \ldots Z_n \sim P$  the empirical bootstrap estimate of the **distribution**  $Q_P$  of  $T_{n,\alpha}$  is found as follows:
  - Estimate P by  $\hat{P}_n$  the empirical distribution of the sample, and, hence,  $Q_P$  by  $Q_{\hat{P}}$
  - Estimate  $Q_{\hat{P}}$  by the empirical distribution of a sample  $T_1^*, \dots T_B^*$  from it

In computational steps this scheme equals:

- Generate B times a sample  $Z_1^*, \ldots Z_n^*$  by resampling with replacement from the initial sample  $Z_1, \ldots Z_n$ .

- Generate for each  $Z^*$ -sample  $T^* = T_{n,\alpha}(Z_1^*, \dots Z_n^*)$ . This yields the bootstrap values  $T_1^*, \dots T_B^*$ .

The bootstrap estimate of the **standard deviation** of  $T_{n,\alpha}$  is found in both schemes by the last step:

- Estimate the sd of  $T_{n,\alpha}$  by the sd of the bootstrap values  $T_1^*, \dots T_B^*$ .
- b. [2 points] Since the data are skewed to the right, the mean (0% trim) is larger than the median (50% trim). Since the data is monotonically skewed, it follows that  $T_{n,\alpha}$  decreases when  $\alpha$  increases. Since the  $\alpha_2$ -trimmed mean is lower than the  $\alpha_1$ -trimmed mean it follows that  $\alpha_1 < \alpha_2$ .
- c. [2 points] The formula for the bootstrap confidence interval is  $[2T T^*_{\lfloor (1-\alpha_1)B}, 2T T^*_{\lfloor \alpha_1 B \rfloor}]$ , which equals [2\*0.7 1.16, 2\*0.7 0.45] = [0.24, 0.95].

#### PART 2

## Question 4 [6 points]

- a. [2 points] Incorrect. The distribution of the KS test statistic need not be normal if the data is from a normal distribution. In general, one uses bootstrap methods to mimic the distribution of a statistic because it is unknown (and certainly not normal).
- b. [2 points] Correct. Hat values only show that an observation is potentially influential, whereas Cook's distances show the real influence.
- c. [2 points] Correct. The a.r.e. of the Wilcoxon signed rank test to the sign test is 3/4 which is lower than 1. So the Wilcoxon signed rank test needs more data to obtain the same power as the sign test. In other words: the sign test has higher power.

## Question 5 [5 points]

- a. [4 points]
  - sign test: not suitable since the sign test is for testing the underlying median of ONE sample.
  - Wilcoxon two sample (rank sum) test: suitable.
  - Kendall's rank correlation test: not suitable, since a rank correlation test is testing dependence in paired samples.
  - Kolmogorov-Smirnov two sample test: suitable
- b. [1 point] Since the Wilcoxon rank sum test is best suited for shift alternatives, it is expected not to have much power in this situation, because the biggest difference between x and y is in scale. The Kolmogorov–Smirnov test is suited for any differences between underlying distributions and is expected to have higher power in this situation.

# Question 6 [7 points]

- a. [3 points] This concerns one multinomial sample of size n=20. The underlying distribution is multinomial  $(20, p_{11}, p_{12}, p_{21}, p_{22})$  with  $p_{11} + p_{12} + p_{21} + p_{22} = 1$ . The null hypothesis of independence between handedness and having voted is  $H_0: p_{ij} = p_{i\cdot}p_{\cdot j}$  with  $p_{i\cdot} = pi1 + p_{i2}$  and  $p_{\cdot j} = p1j + p_{2j}$ .
- b. [3 point] Fisher's exact test uses  $N_{11}$  as test statistic which has under the given  $H_0$  a hypergeometric distribution with  $n=20,\ l=7,$  m=2. Since the observed count in cel (1,1) equals 0 the left p-value equals

 $P(X=0) = \frac{\binom{2}{0}\binom{18}{7}}{\binom{20}{7}} = 0.41.$ 

This one-sided p-value is bigger than any realistic significance level, so  $H_0$  is not rejected. Alternative, one may use  $n=20,\ l=2,\ m=7$  and

 $P(X=0) = \frac{\binom{7}{0}\binom{13}{2}}{\binom{20}{2}} = 0.41.$ 

c. [1 point] The rule of thumb is:  $EN_{ij} > 1$  for all i,j and  $EN_{ij} > 5$  for at least 80% of the cells. Here we have  $EN_{11} = 2 \times 7/20 = 14/20 < 1$  so the rule of thumb does not hold. The chi-square test is therefore not applicable for these data.

# Question 7 [9 points]

a. [3 points] The multiple linear regression model is:

$$Y_i = \beta_0 + \beta_1 x_{i1} + \ldots + \beta_p x_{ip} + e_i$$

with

 $Y_i - i^{th}$ observed response value

 $\beta_0, \dots, \beta_p$  – unknown parameters

 $x_{i1}, \dots, x_{ip}$  — measured explanatory variables for  $i^{th}$  observation

 $e_i$  – error in  $i^{th}$  observation

The assumption on the errors is  $e_i \sim \mathcal{N}(0, \sigma^2)$  i.i.d. with  $\sigma^2 > 0$  the unknown error variance.

- b. [3 points]
  - **linearity** from scatter plots  $(Y, X_j)$  for j = 1, ..., p and added variable plots
  - independence of errors from context
  - **normality of errors** from QQ-plot of residuals  $\hat{e}$  against the normal distribution
  - constant error variance from scatter plots  $(Y, \hat{e}), (\hat{Y}, \hat{e})$
- c. [3 points] The main problem to expect is collinearity. Apart from that there may be some influence points because of outliers in the Armed.Forces values. Remedies for collinearity: scale the design matrix, compute variation inflation factors, condition indices and variance decompositions, and based on this information take a selection of the collinear explanatory variables. Remedies for the influence points: compute hat values and Cook's distances.

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