

**Exam:**  
**Stochastic Gradient Techniques in Optimization and Learning**  
**period 4.2**

**December 2019**

**Problem 1 (15 Credits)**

Let  $J(\theta) = 3 \sin(2\theta + c)$ , for some  $c > 0$ , and  $\theta \in \mathbb{R}$ . For  $\alpha \in (0, 1)$ , we want to find  $\theta^*$  such that

$$J(\theta^*) = \alpha. \tag{1}$$

Let

$$G(\theta) = -(J(\theta) - \alpha).$$

- (a). [5 Credits] Show that the stationary points of  $J(\theta)$  fail to be stable points of  $G(\theta)$ .
- (b). [5 Credits] Under what conditions on  $\alpha$  does  $\theta^*$  become asymptotically stable?
- (c). [5 Credits] Rewrite  $J(\theta^*) = \alpha$  as optimization problem

$$\min_{\theta} \frac{1}{2} (J(\theta) - \alpha)^2.$$

Argue that  $G(\theta)$  is not coercive for this optimization problem, i.e., for tracking the solutions of  $J(\theta^*) = \alpha$ .

**Answer Problem 1:**

- (a) At stationary points the value of  $J(\theta)$  is either -3 or 3. The vector field  $G(\theta)$  moves towards  $3 > \alpha > -3$  and therefore the stationary points are not stable points of  $G(\theta)$ .
- (b)  $\theta^*$  is asymptotically stable if  $G(\theta)$  moves towards  $\theta^*$  when it is in the neighborhood of  $\theta^*$ . This only holds for the chosen  $G$  if  $J(\theta)$  is monotone increasing in the neighborhood of  $\theta^*$ .
- (c)  $J(\theta) = \alpha$  has infinitely many solutions. Moreover, as  $J(\theta)$  is not monotone  $G(\theta)$  moves in the areas where  $J(\theta)$  is monotone decreasing towards a maximum.

## Problem 2 (total 15 Credits)

The gradient-field of a function  $J(\theta)$  is shown in Figure 1. Apply a steepest descent algorithm for

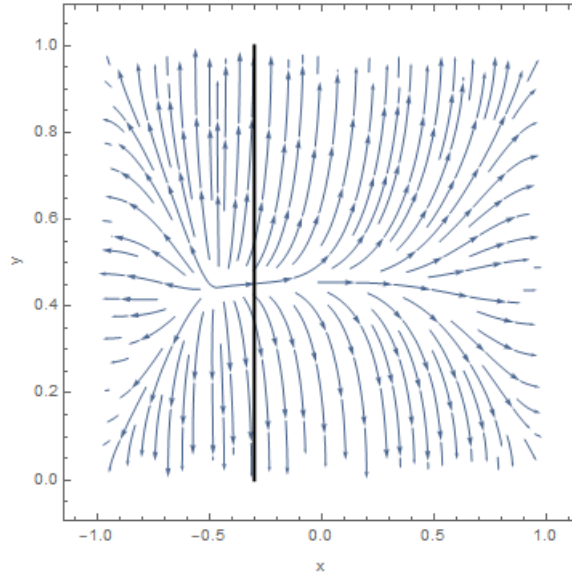


Figure 1: Gradient-field of  $J(\theta)$

finding the minimum of  $J(\theta)$ .

- (a). [5 Credits] Discuss with Figure 1 (where you ignore the black line for this part of the problem) for the ODE

$$\frac{d}{dt}x(t) = -\nabla J(x(t))$$

the nature of point  $(0, 0.5)$  (stable, asymptotically stable, or unstable).

- (b). [5 Credits] Judging from the figure (where you ignore the black line for this part of the problem), is this problem well-posed and is the vector-field coercive?
- (c). [5 Credits] Now suppose we use a descent algorithm for finding the minimum of  $J(\theta)$  on the constraint set given by the black line, i.e., only the points on the black line are admissible. Argue that  $(-0.3, 0.425)$  is an asymptotically stable point for the ODE living on the constraint set.

### Answer Problem 2:

- (a)  $(0, 0.5)$  is not stable (and therefore not asymptotically stable) as some arrows point towards this point while other points away from it.
- (b) Yes, moving in opposite direction of the arrows always leads to approximately  $(-0.5, 0.45)$
- (c) The ODE will move on the black line following the opposite direction of the arrows as much as possible, and this ODE will therefore move to  $(-0.3, 0.425)$ .

### Problem 3 (total 20 Credits)

Consider the algorithm

$$\theta_{n+1} = \theta_n + \epsilon_n Y_n,$$

for finding some optimal solution  $\theta^*$ , for  $\epsilon_n = 1/(n+1)$ , for  $n \in \mathbb{N}$ . Suppose that evaluating  $Y_n$  requires one sample from an underlying process. Suppose your computational budget is sufficient to sample  $N$  samples from the underlying process, and you split your simulation budget to produce  $k$  independent runs of the algorithm yielding  $\theta_n(\omega_i)$ ,  $1 \leq i \leq k$ , for each of the runs, where  $kn = N$ . Running your experiment with  $k = 1000$ ,  $n = 100$ , and  $N = 10^5$  yields the histogram in Figure 2. The black line is the density of the normal distribution fitted to the data.

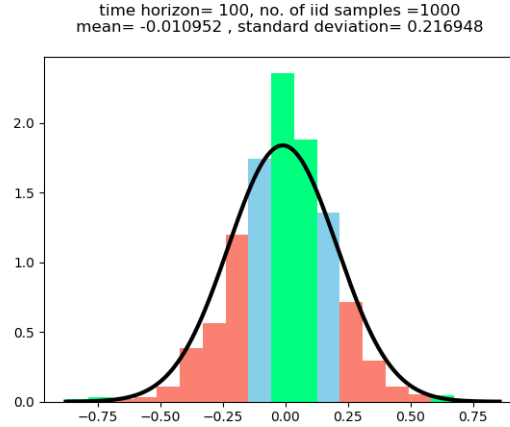


Figure 2: Histogram of  $\theta_{100}$

- [5 Credits] Given the available information, can the claim " $\theta^* = 0.0$ " be rejected at confidence level 0.05? (You may use that  $1/\sqrt{1000} \approx 0.031$ , and  $1.96 \times 0.031 \times 0.2169 \approx 0.0134$ .)
- [5 Credits] Suppose you would run the algorithm with fixed  $\epsilon$  rather than decreasing  $\epsilon_n$ . What would you expect to find for the resulting output? Will the mean value be effected, will the variance be effected? Explain!
- [10 Credits] Keeping the budget fixed, provide a new choice for  $n$  and  $k$  that may improve the statistical properties of  $\theta_n$ .

#### Answer Problem 3:

- 0.0 lies within a 95 % confidence interval and the claim  $\theta^* = 0.0$  cannot be rejected.
- Fixed  $\epsilon$  yields weak convergence, so  $\theta_n$  is a random variable, and for decreasing  $\epsilon$ ,  $\theta_n$  converges a.s. So, fixed  $\epsilon$  will result in  $\theta_n$  with larger variance (decreasing will have eventually zero variance).
- The histogram shows that the normal distribution fit is not very good. This suggests that  $n$  is too small, and it is advisable to increase  $n$  by a factor of, say, 2, and use  $\hat{n} = 2n = 200$ ,  $\hat{k} = 500$ .

**Problem 4 (total 20 Credits)**

Let  $X(\theta)$  be a continuous random variable with a probability cumulative distribution function

$$F_\theta(x) = \begin{cases} 0, & x \leq 0; \\ 1 - e^{-\theta\sqrt{x}}, & x > 0, \end{cases} \quad (2)$$

for  $\theta > 0$ .

(a). [5 Credits] Show that the score function is

$$S(\theta, x) = \frac{1}{\theta} - \sqrt{x}.$$

(b). [10 Credits] Let the parameter space be  $\Theta = [1, 4]$ . Show that

$$\frac{d}{d\theta} \mathbb{E}[h(X(\theta))] = \mathbb{E}[h(X(\theta)) S(\theta, X(\theta))].$$

for  $\theta \in \Theta$  and all polynomially bounded cost functions  $h$ , i.e.,  $h(x) = O(x^k)$  for some positive integer  $k$ .

(c). [5 Credits] Note that (2) is the Weibull distribution with scale parameter  $\theta$ , and with shape parameter  $\alpha = \frac{1}{2}$ . Set  $\theta = 1$ . How would you generate samples from this distribution? Work out a few details.

**Answer Problem 4:**

(a) First, determine the PDF,

$$f_\theta(x) \doteq \frac{\partial}{\partial x} F_\theta(x) = \frac{\theta}{2\sqrt{x}} e^{-\theta\sqrt{x}}, \quad x > 0.$$

The score function is

$$\begin{aligned} S(\theta, x) &\doteq \frac{\frac{\partial}{\partial \theta} f_\theta(x)}{f_\theta(x)} = \frac{\partial}{\partial \theta} \log f_\theta(x) \\ &= \frac{\partial}{\partial \theta} \left( \log \theta - \log(2\sqrt{x}) - \theta\sqrt{x} \right) = \frac{1}{\theta} - \sqrt{x}, \quad x > 0. \end{aligned}$$

(b)

$$\frac{d}{d\theta} \mathbb{E}[h(X(\theta))] = \frac{d}{d\theta} \int_0^\infty h(x) f_\theta(x) dx.$$

A sufficient condition for interchanging  $\frac{d}{d\theta}$  and  $\int$  on  $\Theta = [1, 4]$  is bounded convergence, which holds if

$$\int_0^\infty |h(x)| \sup_{\theta \in \Theta} \left| \frac{\partial}{\partial \theta} f_\theta(x) \right| dx < \infty.$$

Do the calculus,

$$\left| \frac{\partial}{\partial \theta} f_\theta(x) \right| = \left| \frac{1}{2\sqrt{x}} - \frac{\theta}{2} \right| e^{-\theta\sqrt{x}} \leq \left( \frac{1}{\sqrt{x}} + 2 \right) e^{-\sqrt{x}},$$

for  $\theta \in \Theta = [1, 4]$ . Clearly, for any  $k \geq 0$ ,

$$\int_0^\infty x^k \left( \frac{1}{\sqrt{x}} + 2 \right) e^{-\sqrt{x}} dx < \infty.$$

Hence, interchange is allowed for  $h(x) = O(x^k)$ , and results in

$$\begin{aligned} \frac{d}{d\theta} \mathbb{E}[h(X(\theta))] &= \frac{d}{d\theta} \int_0^\infty h(x) f_\theta(x) dx = \int_0^\infty h(x) \frac{d}{d\theta} f_\theta(x) dx \\ &= \int_0^\infty h(x) \frac{\frac{d}{d\theta} f_\theta(x)}{f_\theta(x)} f_\theta(x) dx = \mathbb{E}[h(X(\theta)) S(\theta, X(\theta))]. \end{aligned}$$

(c) Let  $X = X(1)$  be the random variable for  $\theta = 1$ . It has CDF

$$F(x) = 1 - e^{-\sqrt{x}}, \quad x > 0.$$

It holds that  $F(X) = U$ , the uniform random variable on  $(0, 1)$ , thus  $X = F^{-1}(U)$ . In other words, to get a sample  $x$  of  $X$ , it suffices to solve  $F(x) = u$  for any  $u \in (0, 1)$ .

$$\begin{aligned} F(x) = u &\Leftrightarrow 1 - e^{-\sqrt{x}} = u \Leftrightarrow e^{-\sqrt{x}} = 1 - u \\ &\Leftrightarrow -\sqrt{x} = \ln(1 - u) \Leftrightarrow x = (-\ln(1 - u))^2. \end{aligned}$$

Note that because  $1 - U \stackrel{\mathcal{D}}{=} U$ , and  $(-1)^2 = 1$ , you might do

$$x = (\ln(u))^2, \quad u \in (0, 1).$$

**Problem 5 (total 30 Credits)**

- (a). [6 Credits] Let  $\Theta = (a, b) \subset \mathbb{R}$  be a finite interval, and  $f : \Theta \rightarrow \mathbb{R}$  a (real-valued) function on it.
- (i). Give the definition of Lipschitz continuity of  $f$  on  $(a, b)$ .
  - (ii). Suppose that  $f$  is differentiable on  $(a, b)$ ; then give a sufficient condition for Lipschitz continuity that is more easy to check than the definition in (i).
- (b). [9 Credits] Are the following (deterministic) functions Lipschitz continuous? If yes, show by applying (a)-(i) or (a)-(ii); if no, argue that (a)-(i) does not hold.
- (i).  $f(\theta) = \theta e^\theta$  on  $(1, 2)$ .
  - (ii).  $f(\theta) = |\theta|$  on  $(-1, 1)$ .
  - (iii).  $f(\theta) = \log \theta$  on  $(0, 1)$ .
- (c). [15 Credits] Is  $Y(\theta)$  almost surely Lipschitz continuous in the following cases? And if so, is the Lipschitz modulus integrable? Just an answer is not sufficient. Provide an analysis where you proof your claims.
- (i).  $Y(\theta) = X/\theta$  where  $X \stackrel{\mathcal{D}}{\sim} \text{Ex}(1)$  (the exponential distribution with parameter 1), and  $\theta \in \Theta = (1, 2)$ .
  - (ii).  $Y(\theta) = 1/(\theta U)$  where  $U \stackrel{\mathcal{D}}{\sim} U(0, 1)$  (uniform distribution), and  $\theta \in \Theta = (1, 2)$ .
  - (iii).  $Y(\theta) = \sqrt{|X - \theta|}$  where  $X \stackrel{\mathcal{D}}{\sim} \text{Ex}(1)$  (the exponential distribution with parameter 1), and  $\theta \in \Theta = (1, 2)$ .

**Answer Problem 5:**

- (a) (i). There exists  $0 < K < \infty$  such that for any  $\theta_1, \theta_2 \in \Theta$ ,

$$|f(\theta_1) - f(\theta_2)| \leq K|\theta_1 - \theta_2|.$$

- (ii). You may take

$$K = \sup_{\theta \in \Theta} |f'(\theta)|$$

as Lipschitz constant.

- (b) (i). Yes. Apply (a)(ii):

$$\sup_{\theta \in (1, 2)} |f'(\theta)| = \sup_{\theta \in (1, 2)} |1 + \theta|e^\theta = 3e^2 < \infty.$$

- (ii). Yes. Apply (a)(i): w.l.o.g., assume  $-1 < \theta_2 < \theta_1 < 1$ .

$$|f(\theta_1) - f(\theta_2)| = ||\theta_1| - |\theta_2|| = \begin{cases} |\theta_1 - \theta_2|, & -1 < \theta_2 < \theta_1 \leq 0; \\ |\theta_1 - \theta_2|, & 0 \leq \theta_2 < \theta_1 < 1; \\ |\theta_1 - (-\theta_2)| \leq |\theta_1 - \theta_2|, & -1 < \theta_2 < 0 < \theta_1 < 1. \end{cases}$$

Thus Lipschitz constant  $K = 1$ .

(iii). No. Let  $\theta_1 = 0.5$  and  $\theta_2 = 1/n$ , where  $n = 1, 2, \dots$ . Then

$$|\log \theta_1 - \log \theta_2| = \log(n/2),$$

which you cannot bound for all  $n$ . Equivalently,

$$|f'(\theta)| = \frac{1}{\theta} \xrightarrow{\theta \downarrow 0} \infty.$$

(c) (i). Let  $x > 0$  be a random element of  $X$ . Then, the function  $Y(\theta) = x/\theta$  is Lipschitz continuous on  $\Theta = (1, 2)$ . For instance, apply (a)(ii):

$$K(x) = \sup_{\theta \in \Theta} |Y'(\theta)| = \sup_{\theta \in (1,2)} x/\theta^2 = x < \infty.$$

This holds for any  $x > 0$  drawn from the exponential(1) distribution, thus it holds with probability one. The Lipschitz constant becomes the random variable  $K = X$ , for which  $\mathbb{E}[X] = 1 < \infty$ .

(ii). Let  $u \in (0, 1)$  be a random element of  $U$ . Then, the function  $Y(\theta) = 1/(u\theta)$  is Lipschitz continuous on  $\Theta = (1, 2)$ . The Lipschitz constant can be taken to be

$$K(u) = \sup_{\theta \in \Theta} |Y'(\theta)| = \sup_{\theta \in (1,2)} 1/(u\theta^2) = 1/u.$$

This holds for any  $u \in (0, 1)$  drawn from the uniform distribution, thus it holds with probability one. The Lipschitz constant becomes the random variable  $K = 1/U$ , for which

$$\mathbb{E}[1/U] = \int_0^1 \frac{1}{u} du = \infty.$$

(iii). Let  $x > 0$  be a random element of  $X$ , and suppose that  $x \in (1, 2)$ . Then

$$Y(\theta) = \sqrt{|x - \theta|} = \begin{cases} \sqrt{x - \theta}, & 1 < \theta \leq x; \\ \sqrt{\theta - x}, & x \leq \theta < 1. \end{cases}$$

This function is not Lipschitz. Taking  $\theta = x + \epsilon$ , and letting  $\epsilon \downarrow 0$  results in an unbounded derivative

$$Y'(\theta) = Y'(x + \epsilon) = \frac{1}{2\sqrt{\epsilon}}.$$

This holds for all  $x \in (1, 2)$ , and thus with positive probability.

### Bonus Problem (total 20 Credits)

Consider the one-dimensional fitting problem

$$\min_{\theta} \mathbb{E}[(\theta X - X^2)^2]$$

for finding the best "scaling" of  $X$  that produces  $X^2$ .

- (a). [10 Credits] Find an SA for solving this problem.
- (b). [10 Credits] Show that, in general,  $\theta = \mathbb{E}[X]$  is not the correct answer.

#### Answer Bonus Problem:

- (a) Apply IPA, which is straightforward in this case:

$$\frac{d}{d\theta} \mathbb{E}[(\theta X - X^2)^2] = \mathbb{E}[2X(\theta X - X^2)].$$

The negative gradient is coercive for this problem as we deal with minimizing a distance. The SA looks like

$$\theta_{n+1} = \theta_n - \epsilon_n 2X_n(\theta X_n - X_n^2)$$

for  $X_n$  the  $n$ -th observation. To ensure convergence we have to control the variance. As the variance of  $Y_n$  scales in  $\theta^2$ , we need to use a truncation argument.

- (b)

$$\mathbb{E}[(\theta X - X^2)^2] = \theta^2 \mathbb{E}[X^2] - 2\theta \mathbb{E}[X^3] + \mathbb{E}[X^4]$$

Taking derivatives gives,

$$\mathbb{E}'[(\theta X - X^2)^2] = 2\theta \mathbb{E}[X^2] - 2\mathbb{E}[X^3],$$

yielding as stationary point

$$\theta = \frac{\mathbb{E}[X^3]}{\mathbb{E}[X^2]}.$$

As the the second order derivative of  $\mathbb{E}[(\theta X - X^2)^2]$  is positive, this point is a minimum. The fact that we only find one stationary point, show that we have found the global minimum.