

# Exam Applied Stochastic Modeling - Solutions

The solutions are always provisional

December 17, 2018, 8:45 - 11:30 hours

## Exercise 1.

a. It follows directly that the mean service time equals  $\mathbb{E}S = 1$ . The second moment follows from

$$\mathbb{E}S^2 = \int_0^2 \frac{1}{2}u^2 du = \frac{1}{2} \frac{1}{3} u^3 \Big|_{u=0}^2 = \frac{4}{3}.$$

Consequently, the residual service time upon arrival reads

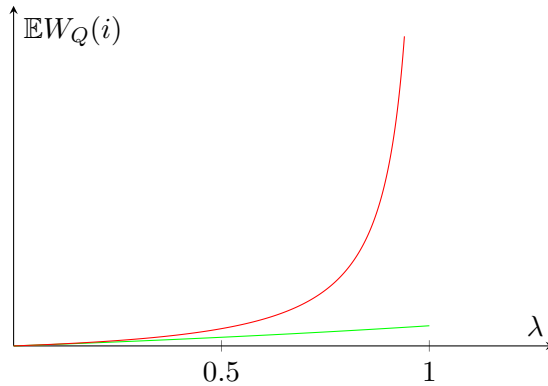
$$\mathbb{E}R = \frac{\lambda \mathbb{E}S^2}{2} = \frac{2}{3}\lambda.$$

Now, using the loads  $\rho_1 = 1/2\lambda \times 1/2 = \lambda/4$ ,  $\rho_2 = 1/2\lambda \times 3/2 = 3\lambda/4$ , and combining the above, we obtain

$$\begin{aligned} \mathbb{E}W_Q(1) &= \frac{\mathbb{E}R}{1 - \rho_1} = \frac{2\lambda/3}{1 - \lambda/4} \\ \mathbb{E}W_Q(2) &= \frac{\mathbb{E}R}{(1 - \rho_1)(1 - \rho_1 - \rho_2)} = \frac{2\lambda/3}{(1 - \lambda/4)(1 - \lambda)} \end{aligned}$$

giving the desired result.

b. See the figure below for a sketch of  $\mathbb{E}W_Q(1)$  (green line) and  $\mathbb{E}W_Q(2)$  (red line). Both functions are increasing and convex.



For  $\lambda \rightarrow 1$ , it holds that  $\mathbb{E}W_Q(1) = 8/9 < \infty$ , whereas  $\mathbb{E}W_Q(2) \rightarrow \infty$  for  $\lambda \rightarrow 1$ . Class 1 is only affected by the load of class 1 ( $\rho_1$ ) and the residual service time; as  $\rho_1$  is strictly smaller than 1 (in fact, at most  $1/4$ ), the waiting time of class 1 remains bounded. Class 2 is affected by the total load  $\rho_1 + \rho_2 = \lambda$ , which converges to 1 such that the total number of customers tends to grow large.

**Exercise 2.**

a. The system is stable for  $\alpha < \mu$ .

b. Let  $X(t)$  denote the number of customers at time  $t$ . The transition diagram of the birth-and-death process  $X(t)_{t \geq 0}$  is given in Figure 1. The distribution of the number of customers in the system follows from the balance equations (for sets):  $4\pi(0) = \mu\pi(1)$  and  $\alpha\pi(i-1) = \mu\pi(i)$ , for  $i = 2, 3, \dots$ . The first equation yields  $\pi(1) = \frac{4}{\mu}\pi(0)$ , whereas the combination gives

$$\pi(i) = \left(\frac{\alpha}{\mu}\right)^{i-1} \pi(1) = \left(\frac{\alpha}{\mu}\right)^{i-1} \frac{4}{\mu} \pi(0), \quad i = 1, 2, \dots$$

Using normalization, we obtain that

$$\pi(0) = \left[1 + \sum_{i=1}^{\infty} \left(\frac{\alpha}{\mu}\right)^{i-1} \frac{4}{\mu}\right]^{-1} = \left[1 + \frac{4}{\mu - \alpha}\right]^{-1} = \frac{\mu - \alpha}{\mu - \alpha + 4}.$$

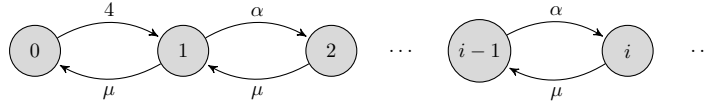


Figure 1: State diagram Exercise 2b.

c. If  $\alpha = 0$ , then the process is similar to that of an M/M/1/1 queue. Regeneration epochs can be the moments at which the queue becomes empty (just after the service completion)<sup>1</sup>. We have that the expected cycle length is  $\mathbb{E}T = \frac{1}{4} + \frac{1}{\mu}$ . The costs per cycle is the idle time per cycle, hence  $\mathbb{E}[\text{costs per cycle}] = \frac{1}{4}$ . Using the renewal reward theorem, we find

$$\mathbb{P}(\text{idle}) = \frac{\mathbb{E}[\text{costs per cycle}]}{\mathbb{E}T} = \frac{1/4}{1/4 + 1/\mu} = \frac{\mu}{4 + \mu}.$$

d. First note that we can use the same regeneration epochs as in part c. Also, we still have that  $\mathbb{E}[\text{costs per cycle}] = \frac{1}{4}$ . It remains to determine the expected cycle length<sup>2</sup>, which can be found by conditioning on the idle time:

$$\begin{aligned} \mathbb{E}T &= \int_{u=0}^t \left(u + \frac{1}{\mu}\right) \times 4e^{-4u} du + \int_{u=t}^{\infty} \left(u + \frac{1}{2\mu}\right) \times 4e^{-4u} du \\ &= \int_{u=0}^{\infty} u \times 4e^{-4u} du + \frac{1}{\mu} (1 - e^{-4t}) + \frac{1}{2\mu} e^{-4t} = \frac{1}{4} + \frac{1}{2\mu} (2 - e^{-4t}), \end{aligned}$$

since the integral on the second line equals  $1/4$ . Hence, by the renewal reward theorem again, we have

$$\mathbb{P}(\text{idle}) = \frac{\mathbb{E}[\text{costs per cycle}]}{\mathbb{E}T} = \frac{1/4}{1/4 + \frac{1}{2\mu} (2 - e^{-4t})}.$$

**Exercise 3.**

a. Let  $N(t)_{t \geq 0}$  denote the arrival process of customers (i.e. a Poisson process with rate 10

<sup>1</sup>There are many other regeneration epochs possible in this case.

<sup>2</sup>An alternative is to write  $\mathbb{E}T = \mathbb{E}A + \mathbb{P}(A \leq t) \frac{1}{\mu} + \mathbb{P}(A > t) \frac{1}{2\mu}$ , with  $A$  denoting the interarrival time.

starting at time 0). We apply thinning of Poisson processes as follows: for an arrival at time  $t \in [0, \tau]$ , it is considered to be of type 1 if the service time is larger than  $\tau - t$  (denoted as process  $N_1(t)_{t \in [0, \tau]}$ ) and of type 2 otherwise. Then, for  $t \in [0, \tau]$ ,  $N_1(t)_{t \in [0, \tau]}$  is a Poisson process with rate  $10e^{-(\tau-t)}$ .

Now, the number of customers present at time  $\tau$  equals  $N_1(0, \tau)$ , with  $N_1(s, t)$  the number of type 1 arrivals during  $[s, t]$ . Note that  $N_1(0, \tau)$  is a Poisson random variable with rate

$$m(\tau) = \int_0^\tau 10e^{-(\tau-t)} dt = 10 \int_0^\tau e^{-t} dt = 10(1 - e^{-\tau}),$$

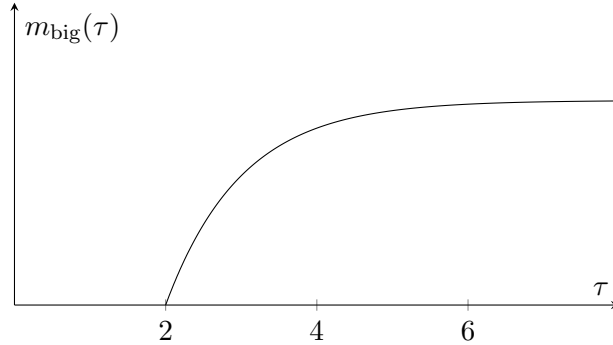
giving the desired result.

b. Observe that a customer arriving at time  $t \in [0, \tau]$  is still present at time  $\tau$  if the service time exceeds  $\tau - t$ . Hence, we may use the same splitting as in part a. If a customer at time  $\tau$  is big, it must have arrived before time  $\tau - 2$ . Thus, the number of big customers at time  $\tau$  equals  $N_1(0, \tau - 2)$ , for  $\tau \geq 2$ . Note that  $N_1(0, \tau - 2)$  is a Poisson random variable with rate

$$m_{\text{big}}(\tau) = \int_0^{\tau-2} 10e^{-(\tau-t)} dt = 10 e^{-\tau} e^{+t} \Big|_{t=0}^{\tau-2} = 10(e^{-2} - e^{-\tau}),$$

giving the desired result.

c. Please see the figure below for a sketch of  $m_{\text{big}}(\tau)$ . The figure displays that the queue gradually increases until it reaches its equilibrium, i.e., reflecting the startup of the system.



#### Exercise 4.

a. The routing equations are  $\gamma_1 = \lambda + p_1\gamma_1$  and  $\gamma_2 = p_2\gamma_1$ . This gives  $\gamma_1 = \lambda/(1-p_1)$  and  $\gamma_2 = p_2\lambda/(1-p_1)$ . The system is stable if  $\gamma_1/3 < 1$  and  $\gamma_2/3 < 1$ . As  $\gamma_2 < \gamma_1$ , we need that  $\lambda/(1-p_1) < 3$  for the system to be stable.

This is a Jackson network and the stationary distribution is thus of product form:

$$\begin{aligned} \pi(n_1, n_2) &= \left(1 - \frac{\gamma_1}{3}\right) \left(\frac{\gamma_1}{3}\right)^{n_1} \left(1 - \frac{\gamma_2}{3}\right) \left(\frac{\gamma_2}{3}\right)^{n_2} \\ &= \left(1 - \frac{\lambda}{3(1-p_1)}\right) \left(\frac{\lambda}{3(1-p_1)}\right)^{n_1} \left(1 - \frac{p_2\lambda}{3(1-p_1)}\right) \left(\frac{p_2\lambda}{3(1-p_1)}\right)^{n_2}. \end{aligned}$$

b. Let  $X_i(t)$  denote the number of customers at station  $i$  at time  $t$ . The transition diagram of the Markov process  $(X_1(t), X_2(t))_{t \geq 0}$  is depicted in Figure 2. The balance equations are, for  $n_2 \geq 1$ ,

$$\begin{aligned} (\lambda + 6)\pi(n_1, n_2) &= \lambda\pi(n_1 - 1, n_2) + 3\pi(n_1 + 1, n_2 - 1) + 3\pi(n_1, n_2 + 1), \quad n_1 \geq 1 \\ (\lambda + 6)\pi(0, n_2) &= 6\pi(0, n_2 + 1) + 3\pi(1, n_2 - 1). \end{aligned}$$

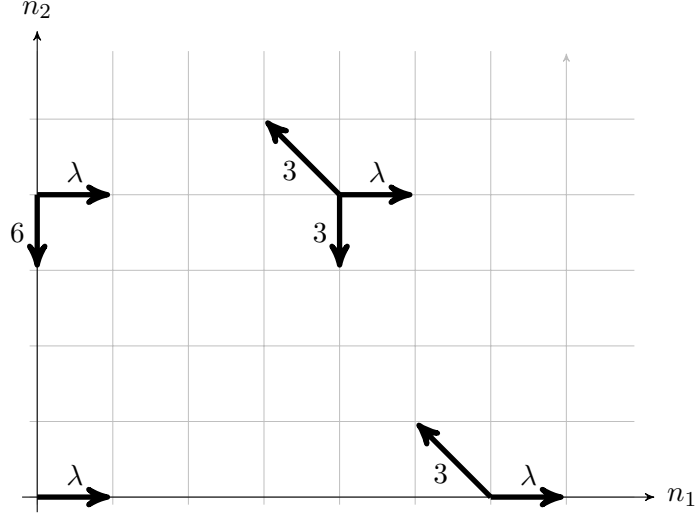


Figure 2: State diagram for Exercise 4b. Only outgoing transitions are shown.

**Exercise 5.**

a. The three terms of  $P(S)$  correspond to the three different sales options: (i)  $K - S$  is the deterministic number of items sold to the trader, (ii)  $\mathbb{E} \min(D, S)$  is the expected sales that they try to sell themselves, and (iii)  $\mathbb{E}(S - D)^+$  are the expected number of unsold items that go to the auction.

b. Assume for now that the demand is continuous and let  $F_D(\cdot)$  denote the distribution function. Note that  $\mathbb{E} \min(D, S) = S - \mathbb{E}(S - D)^+$ . Hence, the expected income can also be written as

$$P(S) = p_2 K - (p_2 - p_1)S + (v - p_1)\mathbb{E}(S - D)^+.$$

Taking derivatives with respect to  $S$  yields  $P'(S) = p_1 - p_2 + (v - p_1)F_D(S)$ , since  $\frac{d}{dS} \mathbb{E}(S - D)^+ = F_D(S)$ . Setting  $P'(S) = 0$  (and noting this provides the unique maximum) gives  $F_D(S) = (p_1 - p_2)/(p_1 - v)$ , or

$$S^* = F_D^{-1} \left( \frac{p_1 - p_2}{p_1 - v} \right).$$

If the demand is discrete, we may use marginal arguments. If the organization tries to sell the  $S$ th item themselves, the expected income is  $p_1(1 - F_D(S)) + vF_D(S) = p_1 - (p_1 - v)F_D(S)$ . If this item is sold to the trader, the income is  $p_2$ . Hence, the organization should try to sell themselves as long as  $p_1 - (p_1 - v)F_D(S) \geq p_2$ , i.e.  $S^*$  is the largest integer that satisfies this equation.

c. If  $v > p_2$  it is always more profitable to try to sell themselves, as the value at the auction is also larger than the value obtained from the trader. Hence,  $S^* = K$  in this case. Observe that the  $(p_1 - p_2)/(p_1 - v) > 1$  (if  $v \leq p_1$ ), so that the inverse in that point is not well defined.